Denoising using wavelets

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Threshold functions







• Threshold functions $s_{\lambda}(t)$

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- are used to suppress parts of a signal with very low amplitudes and (usually) high frequencies ("noise")
- examples are

$$\begin{aligned} \{ \mathsf{Hard}, \lambda \} & \begin{cases} 0 & |x| \leq \lambda \\ x & |x| > \lambda \end{cases} \\ \{ \mathsf{Soft}, \lambda \} & \begin{cases} 0 & |x| \leq \lambda \\ \mathsf{sgn}(x)(|x| - \lambda) & |x| > \lambda \end{cases} \\ \{ \mathsf{PiecewiseGarrote}, \lambda \} & \begin{cases} 0 & |x| \leq \lambda \\ x - \frac{\lambda^2}{x} & |x| > \lambda \end{cases} \\ \\ \mathsf{SmoothGarrote}, \lambda, n \} & \frac{x^{2n+1}}{x^{2n} + \lambda^{2n}} \\ \{ \mathsf{Hyperbola}, \lambda \} & \begin{cases} 0 & |x| \leq \lambda \\ \mathsf{sgn}(x)\sqrt{x^2 - \lambda^2} & |x| > \lambda \end{cases} \end{aligned}$$

.



Figure: Examples of threshold functions $s_{\lambda}(t)$

- Setting and strategy
 - "true" signal : $\boldsymbol{v} = (v_1, v_2, \dots, v_N)$
 - noise vector: $\boldsymbol{\varepsilon} = (\varepsilon_1, \varepsilon_2, \dots, \varepsilon_N)$
 - noised signal $oldsymbol{y} = (y_1, y_2, \dots, y_N) = oldsymbol{v} + arepsilon$
 - wavelet filtering (orthogonal transform!)

$$\boldsymbol{y} \stackrel{WT}{\longmapsto} \boldsymbol{z} = (\boldsymbol{a}, \boldsymbol{d}) = (H\boldsymbol{y}, G\boldsymbol{y})$$

• applying thresholding with $s_\lambda(t)$ to the high-pass component

$$m{z}\longmapsto \widehat{m{z}}=(m{a},\widehat{m{d}})$$
 with $\widehat{m{d}}=s_{\lambda}(m{d})$

• inverse wavelet transform

$$\widehat{\boldsymbol{z}} \stackrel{WT^{-1}}{\longmapsto} \widehat{\boldsymbol{v}} = H^{\dagger} \boldsymbol{a} + G^{\dagger} \widehat{\boldsymbol{d}}$$

- Heuristic considerations
 - Noise modelled as *Gaussian white noise* with noise level σ : $\varepsilon = (\varepsilon_1, \varepsilon_2, \dots, \varepsilon_N)$ generated by independent and identically $\mathcal{N}(0, \sigma^2)$ -distributed random variables
 - For a vector $\varepsilon = (\varepsilon_1, \ldots, \varepsilon_N)$ of independent, $\mathcal{N}(0, \sigma^2)$ -distributed random variables ε_i and an orthogonal $(N \times N)$ matrix U the components γ_i of the vector

$$\gamma = (\gamma_1, \ldots, \gamma_N) = U \varepsilon,$$

are again independent $\mathcal{N}(\mathbf{0},\sigma^2)$ -distributed random variables

• Due to orthogonality of the wavelet transform the transformed noise term $WT(\varepsilon)$ in

$$\mathbf{y} = \mathbf{v} + \mathbf{\varepsilon} \longmapsto WT(\mathbf{y}) = WT(\mathbf{v}) + WT(\mathbf{\varepsilon})$$

is still characterized by being white noise with noise level $\boldsymbol{\sigma}$

- Heuristic considerations (contd.)
 - In wavelet transformations most energy goes into the approximation (low-pass) component **a**
 - Noise of high frequency goes into the detail (high-pass) component ${m d}$
 - \implies the detail component mainly (but not exclusivlely) consists of noise (detail coefficients $\leq \sigma$) that is where to attack!
 - The problem: the noise level σ is not known and has to be estimated from the data to be denoised themselves
 - How to choose λ (depending on the estimate for σ) ?
 - Measure of quality for denoising: mean squared error (MSE)

$$\mathsf{E}\left[\|\mathbf{v}-\widehat{\mathbf{v}}\|^{2}\right] = \mathsf{E}\left[\sum_{1 \leq j \leq N}(v_{j}-\widehat{v_{j}})^{2}\right]$$

- $\boldsymbol{v} = (v_1, \ldots, v_N) \in \mathbb{R}^N$
- $\boldsymbol{\varepsilon} = (\varepsilon_1, \dots, \varepsilon_N)$ white noise with variance σ^2 , $\boldsymbol{y} = \boldsymbol{v} + \boldsymbol{\varepsilon}$
- $\widehat{\boldsymbol{v}} = (\widehat{v}_1, \dots, \widehat{v}_N)$ estimate for \boldsymbol{v} with $A \subseteq \{1, 2, \dots, N\}$ and

$$\widehat{v}_j = \begin{cases} y_j & \text{if } j \in A \\ 0 & \text{if } j \notin A \end{cases}$$

In this case the MSE equals

$$\mathsf{E}\left[\|\boldsymbol{v}-\widehat{\boldsymbol{v}}\|^2\right] = \mathsf{E}\left[\sum_{1 \leq j \leq N} (v_j - \widehat{v}_j)^2\right] = \sum_{j \in A} \mathsf{E}\left[\varepsilon_j^2\right] + \sum_{j \notin A} \mathsf{E}\left[v_j^2\right]$$

and this is minimized by setting $j \in A \iff v_j^2 > \sigma^2$ • so that the ideal MSE is

$$\mathsf{E}\left[\|\boldsymbol{v}-\widehat{\boldsymbol{v}}\|^{2}\right] = \sum_{1 \leq j \leq N} \min(v_{j}^{2}, \sigma^{2})$$

- VISUSHRINK :
 - is Wavelet shrinkage with

 $\lambda = \lambda^{\text{univ}} = \sigma \cdot \sqrt{2 \ln N}$ ("universal tolerance")

• Theorem [DONOHO-JOHNSTONE, 1995] For $\mathbf{v} \in \mathbb{R}^N$ and $\boldsymbol{\varepsilon} = (\varepsilon_1, \dots, \varepsilon_N)$ white noise with noise level σ , for $\hat{\mathbf{v}} = s_{\lambda}(\mathbf{v} + \boldsymbol{\varepsilon})$ (soft threshold) one gets

$$\mathsf{E}\left[\|\boldsymbol{v}-\widehat{\boldsymbol{v}}\|^{2}\right] \leq (2\ln N+1) \cdot \left(\sigma^{2} + \sum_{1 \leq j \leq N} \min(v_{j}^{2}, \sigma^{2})\right)$$

- σ will be estimated (on the first high-pass component!) using the mean absolute deviation (MAD):
 - $w = (w_1, ..., w_N)$
 - $\widetilde{w} =$ Median of w

•
$$\mathbf{v} = (|w_1 - \widetilde{\mathbf{w}}|, \dots, |w_N - \widetilde{\mathbf{w}}|)$$

- MAD(w) = Median of $v = \tilde{v}$
- Theorem [HAMPEL, 1974]

$$MAD(w) \approx 0.6745 \cdot \sigma$$

- SURE method (STEINS <u>unbiased</u> risk estimator, 1981)
- Goal: choice of the λ parameter for soft-shrinking methods

$$s_\lambda(x) = egin{cases} x-\lambda & ext{if } x>\lambda \ 0 & ext{if } |x|\leq\lambda \ x+\lambda & ext{if } x<-\lambda \end{cases}$$

C.M. STEIN, Estimation of the mean of a multivariate normal distribution, Ann. Stat. 1981.
 D. DONOHO, I. JOHNSTONE, Adapting to unknown smoothness via wavelet shrinkage, J. Amer. Stat. Assoc. 1995.
 P. VAN FLEET, Discrete Wavelet Transformations, Wiley, 2008 (ch. 9).

Lemma:

For a $\mathcal{N}(0, \sigma^2)$ -distributed random variable ε , any $z \in \mathbb{R}$ and any piecewise differentiable function $g : \mathbb{R} \to \mathbb{R}$ one has

$$\mathsf{E}\left[\varepsilon \cdot g(z+\varepsilon)\right] = \sigma^2 \cdot \mathsf{E}\left[g'(z+\varepsilon)\right]$$

• This follows from partial integration:

$$E[\varepsilon \cdot g(z+\varepsilon)] = \frac{1}{\sqrt{2\pi\sigma^2}} \int x \cdot g(z+x) \cdot e^{-\frac{x^2}{2\sigma^2}} dx$$
$$= \frac{-1}{\sqrt{2\pi\sigma^2}} \int \sigma^2 \cdot g(y) \cdot \frac{d}{dy} e^{-\frac{(y-z)^2}{2\sigma^2}} dy$$
$$= \sigma^2 \cdot \frac{1}{\sqrt{2\pi\sigma^2}} \int \frac{d}{dy} g(y) \cdot e^{-\frac{(y-z)^2}{2\sigma^2}} dy$$
$$= \sigma^2 \cdot E[g'(z+\varepsilon)]$$

• Consequence:

With z, ε, g as in the Lemma, one gets for the MSE of the estimate $\hat{z} = w + g(w)$ of the random variable $w = z + \varepsilon$:

$$\mathsf{E}\left[(\widehat{z} - z)^2\right] = \mathsf{E}\left[(\varepsilon + g(z + \varepsilon))^2\right] \\ = \mathsf{E}\left[(\varepsilon^2 + 2\varepsilon \cdot g(z + \varepsilon) + g(z + \varepsilon)^2\right] \\ = \mathsf{E}\left[\sigma^2 + 2\sigma^2 \cdot g'(w) + g(w)^2\right]$$

• Special case: soft-shrinking with threshold value λ

• The function is

$$g(z) = egin{cases} -z & ext{if } |z| < \lambda \ -\lambda \operatorname{sgn}(z) & ext{if } |z| \geq \lambda \end{cases}$$

and thus

$$rac{d}{dz}g(z) = egin{cases} -1 & ext{if } |z| < \lambda \ 0 & ext{if } |z| \geq \lambda \end{cases}$$

• Therefore

$$\sigma^{2} + 2 \sigma^{2} g'(w) + g(w)^{2} = \begin{cases} w^{2} - \sigma^{2} & \text{if } |w| < \lambda \\ \sigma^{2} + \lambda^{2} & \text{if } |w] \geq \lambda \end{cases}$$

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• General situation:

• Theorem (STEIN)

With the notions just introduced, the MSE can be written as

$$\mathsf{E}\left[\|\widehat{\boldsymbol{z}} - \boldsymbol{z}\|^2\right] = \mathsf{E}\left[\sum_{1 \le j \le N} \left(\sigma_j^2 + 2\sigma_j^2 \frac{\partial}{\partial w_j} g_j(\boldsymbol{w}) + g_j(\boldsymbol{w})^2\right)\right]$$
$$= \|\boldsymbol{\sigma}\|^2 + 2\sum_{1 \le j \le N} \sigma_j^2 \,\mathsf{E}\left[\frac{\partial}{\partial w_j} g_j(\boldsymbol{w})\right] + \mathsf{E}\left[\|\boldsymbol{g}(\boldsymbol{w})\|^2\right]$$

• Special case: soft-shrinking with threshold value λ

$$\mathsf{E}\left[\|\widehat{\boldsymbol{z}} - \boldsymbol{z}\|^2\right] = \mathsf{E}\left[\sum_{1 \le j \le N} \left(w_j^2 - \sigma_j^2 + \left(2\sigma_j^2 - w_j^2 + \lambda^2\right)\chi_{|w_j| \ge \lambda}\right)\right]$$
$$= \mathsf{E}\left[\|\boldsymbol{w} - \boldsymbol{\sigma}\|^2\right] + \mathsf{E}\left[\sum_{1 \le j \le N} \left(2\sigma_j^2 - w_j^2 + \lambda^2\right)\chi_{|w_j| \ge \lambda}\right]$$

The left summand is independent of λ.
 Minimizing the MSE can be achieved by choosing λ depending on the sample vector w so that the integrand in the second summand is minimal!

$$f(\lambda) = \sum_{1 \le j \le N} \left(2\sigma_j^2 - w_j^2 + \lambda^2 \right) \chi_{|w_j| \ge \lambda}$$

• Instead of dealing with $f(\lambda)$ it is more convenient to consider

$$\widetilde{f}(\lambda) = \sum_{1 \leq j \leq N} \left(2\sigma_j^2 - w_j^2 + \lambda^2
ight) \chi_{|w_j| > \lambda}$$

which changes nothing as far as the expectation is concerned

• Assume that the components of the vector $\boldsymbol{w} = (w_1, \dots, w_N)$ are ordered by increasing absolute value

$$|w_0|=0\leq |w_1|\leq |w_2|\leq \cdots \leq |w_N|$$

The function *f*(λ) is continuous from the right for λ ∈ ℝ₊, and if |w_j| < |w_{j+1}| holds, then *f*(λ) is strictly increasing on the half-open interval [|w_j|, |w_{j+1}|), so that it takes its minimum at |w_j|:

$$\widetilde{f}(|w_j|) = (N-j)w_j^2 + \sum_{j+1 \leq k \leq N} (2\sigma_k^2 - w_k^2)$$



Figure: Example for the computation of the minimum value of $f(\lambda)$ for the sequence {0.25, 0.36, 0.41, 0.88, 1.37, 1.48, 1.82, 1.91, 2.3}

From this one gets

$$\min_{\lambda \in \mathbb{R}_+} \widetilde{f}(\lambda) = \min_{0 \le j \le N} \widetilde{f}(|w_j|) \qquad \lambda^{sure} := \operatorname{argmin}_{|w_j|} \widetilde{f}(|w_j|)$$

• The definition of \tilde{f} yields a (downward) recursion

$$\widetilde{f}(|w_j|) = \widetilde{f}(|w_{j+1}|) + 2\sigma_{j+1}^2 + (N-j)(w_j^2 - w_{j+1}^2)$$

which starting from

$$\widetilde{f}(w_N) = 0$$

gives a fast computation of the minimum!

• The usual assumption $\sigma_1 = \ldots = \sigma_N$ somewhat simplifies the formulas and the computations

SURE in the context of the wavelet transform

Consider a one-level WT

$$oldsymbol{z}\longmapstooldsymbol{y}=oldsymbol{z}+arepsilon \stackrel{W\mathcal{T}}{\longmapsto}(oldsymbol{a},oldsymbol{d})\longmapsto(oldsymbol{a},\widehat{oldsymbol{d}})\stackrel{W\mathcal{T}^{-1}}{\longmapsto}\widehat{oldsymbol{z}}$$

where

$$\boldsymbol{d} = \boldsymbol{G}(\boldsymbol{z} + \boldsymbol{\varepsilon}) = \boldsymbol{G}\boldsymbol{z} + \boldsymbol{G}\boldsymbol{\varepsilon} \stackrel{\boldsymbol{s}_{\lambda}}{\longmapsto} \widehat{\boldsymbol{d}}$$

Note: the high-pass component $G\varepsilon$ has the same noise as ε

- σ must be estimated beforehand (as in the VISUSHRINK method)
- One has $\lambda^{sure} \leq \lambda^{univ}$. Recommendation: If y is sparse, it is better to use λ^{univ} instead of λ^{sure} , based on the criterion

$$\frac{1}{N}\sum_{1\leq j\leq n}(y_j^2-\sigma^2)\leq \frac{3}{2\sqrt{N}}\log_2(N)$$

(DONOHOE, JOHNSTONE)